

# Reputation and competition: evidence from the credit rating industry

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Draft Date: November 3, 2008

**Abstract.** Fair and accurate credit ratings arguably play an important role in the financial system. In an environment absent free entry of rating agencies, the provision of quality ratings is at least partially sustained by the reputational concerns of the rating agencies. The economically significant entry of a third agency into a market that was previously best described as a duopoly provides a unique experiment to examine the effect of increased competition on the disciplining effects of reputation. Using a variety of data sources, we find that competition leads to more issuer-friendly and less informative ratings. First, the credit ratings issues by the two incumbent agencies increased toward good ratings. Second, the correlation between bond yields and ratings fell. And lastly, negative stock price responses to announced rating downgrades are larger in absolute value (a downgrade in this weaker ratings environment is even worse news). Ultimately, our findings are consistent with models that suggest competition can impede the reputational mechanism.

*JEL: C7, D83, G14*

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## 1. Introduction

A credit rating is an assessment of the credit worthiness of a corporation or security, most often based on the history of borrowing and repayment for the issuer, its underlying assets, its outstanding liabilities and its overall business performance. These credit ratings fulfill a key function of information transmission in debt markets.<sup>1</sup> Issuers seek ratings for a number of reasons, including to improve the marketability or pricing of their financial obligations, to improve the trust of their business counterparties or because they wish to sell securities to investors with preferences over ratings. Many investors rely on ratings for investment decisions. The amount of capital required for banks and insurance companies who own securities varies with the credit rating. There are also regulatory constraints forcing some investors (e.g., insurance companies and Savings & Loans) to only hold debt securities of investment grade (i.e., with a rating of BBB or better). For these reasons, ratings are considered important by issuers and investors alike.<sup>2</sup>

The provision of accurate ratings is made more complicated by this peculiar market structure of the ratings industry. Ratings of firms and of particular security issues are produced by rating agencies, such as Moody's, Standard & Poor's (S&P) and Fitch Ratings. Once issued, ratings are made publicly available and disseminated for free. Agencies charge the firms they rate (or whose securities they rate) for the work, but the users of ratings, such as investors, use them for free.<sup>3</sup> Users of ratings, such as investors, desire accurate ratings. However, rating agencies' revenues come from fees paid by issuers who themselves or their securities are rated. These issuers likely prefer favorable ratings, and not necessarily fully truthful ones. This

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<sup>1</sup> The majority of ratings of corporate securities relate to corporate bonds. Corporate securities other than bonds, such as preferred stock, are frequently rated as well, and government bonds (at the municipal, state and federal levels) and structured financial products (such as CMOs, CDOs, etc.) are also rated. See Table 1B for an overview of ratings categories.

<sup>2</sup> See Graham and Harvey (2001) for a survey of financial executives' attitudes toward credit ratings, Campbell and Taksler (2003) for recent evidence on the effect of ratings on corporate bond prices, and Tang (2006) regarding the information transmission of ratings. Kisgen (2006) shows how firm capital structure decisions are affected by credit rating considerations.

<sup>3</sup> Early on, rating agencies tried an alternative revenue model that charged users of ratings. This model suffers from being very dependent on the enforcement of contractual limits to how customers can share ratings information they receive. As pointed out by White (2002), the change from user-paid to issuer-paid ratings as the dominant model "in the early 1970s coincides with the spread of low-cost photo-copying".

disagreement leads to an agency problem. The desire of raters to please paying customers may thus potentially lead to compromised quality of credit ratings. It has been suggested that the key feature that keeps this tension in check is the raters' concern for their reputations for issuing honest and accurate ratings (see e.g. Cantor and Packer (1994) and Smith and Ingo (2002)). Rating agencies' reputations are indeed considered critical by industry observers. According to a Bear Stearns & Co equity analyst in June 2007, S&P claimed that "reputation is more important than revenues". Bloomberg news cites Moody's CEO Raymond McDaniel stating that "we are in a business where reputational capital is more important".<sup>4</sup> Former executive VP of Moody's Thomas McGuire stated in 1995 that: "what's driving us is primarily the issue of preserving our track record. That's our bread and butter."<sup>5</sup>

A theoretical literature that begins with Klein and Leffler (1983) argues that the formation of reputations can help support quality provision in markets where information problems would otherwise preclude it (see also e.g. Shapiro 1983, and Cooper and Ross 1984, Diamon 1989, Mailath and Samuelson 2001, Bar-Isaac 2005, and Bar-Isaac and Tadelis forthcoming). In the standard setting, users care about output quality. However, they can only assess the quality of a seller's product after using it, so they must make purchase decisions without this information. The provision of high quality under such imperfect information is not an equilibrium in a one-shot model, but if interactions are repeated, high quality may be sustainable in equilibrium. In Klein and Leffler's multi-period model, producers' quality history is common knowledge and consumers can condition their decisions on that history. In a reputational equilibrium, sellers are induced to provide high quality (at a cost) when the value of expected future rents associated with a maintained reputation exceeds the temporary profit gains from delivering lower quality goods. Something akin to this mechanism may be at work in the ratings industry.

In Klein and Leffler's setting, the building and maintenance of reputation is likely to be heavily affected by competition, and usually for the worse. Competition will reduce the effectiveness of the reputational mechanism for two reasons. First, reputations are only valuable if there are future producer rents. Since competition typically reduces rents, the

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<sup>4</sup> Bloomberg News March-11-2008, "Moody's, S&P Defer Cuts on AAA Subprime, Hiding Loss".

<sup>5</sup> Institutional investor, 10-1995, "Ratings Trouble".

incentive for maintaining a reputation is correspondingly reduced by competition. Second, if the demand elasticity facing individual sellers is higher in a competitive market, the temptation to reduce prices or otherwise attract business may be stronger, again undermining the value of preserving a reputation for high-quality ratings in the hope of garnering future rents.

On the other hand, competition may enhance the effectiveness of the reputational mechanism if the existence of competitive choice is required to make the loss of reputation a real threat. In Hörner's (2002) model, "*competition endogenously generates the outside option inducing disappointed consumers to leave the firm*". Only when consumers have a choice of supplier does the loss of reputation lead to lost business. Hence, Hörner's theory provides an alternative prediction about the effect of increased competition on the reputational mechanism: the quality of ratings may increase with competition (at least over some range).

Both sets of theoretical arguments plausibly apply to the credit rating industry. However, of the two positions, the argument that ratings would be of better quality if there was more competition has been raised most often in policy debates. For example, Paul Schott Stevens, the President of the Investment Company Institute, stated "I firmly believe that robust competition for the credit rating industry is the best way to promote the continued integrity and reliability of their ratings" in testimony for a US Senate Committee on Banking, Housing, and Urban Affairs.<sup>6</sup> Thus, our analysis of this issue is clearly timely.

The credit ratings industry provides a natural environment for studying the effect of competition on reputation and its ability to mitigate potential agency problems.<sup>7</sup> The environment differs from the standard setting where there are two parties: producers who produce a good of ex ante unobservable quality and consumers who must decide whether to buy a product where sellers are evaluated by their reputations. In the credit ratings industry,

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<sup>6</sup> See <http://www.financial-planning.com/asset/article/527499/fund-industry-group-calls-more-credit.html>

<sup>7</sup> Other industries where reputations have been studied empirically include equity analysts (Chevalier Ellison (1999), Hong Kubik (2003) and Hong Kacperczyk (2008)), auto mechanics (Hubbard (2002)), online trading (Cabral Hortaçsu (2006)) and restaurants (Jin Leslie (2003, 2008)). Using a natural experiment (a merger), Hong and Kacperczyk (2008) find that competition reduced biases for earnings analysts (i.e. probably improved quality). Among the many interesting differences, equity analysts face quicker feedback (earning forecasts are usually for less than a year), and have personal (as opposed to firm) reputations.

there are three parties: ratings agencies, firms, and investors. The rating agencies have a reputation for the quality of their ratings. Investors attempt to determine the value of securities using these ratings. Firms must then choose which rating agencies (if any) to use, based on whether investors will assign value to them. Also, regulation may require investors to pay particular attention to ratings (e.g., many investors – pension funds, insurance companies, etc. – can only consider holding securities with investment grade ratings). Owing to these factors, the credit ratings industry is a slightly more complicated setting than that used in standard models. However, since the fundamental research question is the same (whether reputations can motivate the production of temporarily unobservable high quality services), it seems natural that many predictions should apply. A few conditions are obvious. First, theories of reputation will only apply to the ratings industry if investors and issuers (firms) agree about the reputation established by raters. This seems plausible since corporate default, the key event for the ex post assessment of ratings quality, is publicly observable. Other information regarding security payoffs is also typically common knowledge once the payoffs are realized. Second, the provision of quality must be costly to producers. Presumably, informative ratings are expensive to produce because they require the input of significant skilled labor and require upfront investment in experience and industry knowledge, whereas low quality ratings are cheaper to produce. Perhaps the largest potential cost to providing honest and accurate ratings is the potential forgone revenue from unhappy issuers.<sup>8</sup>

Until the late 1990s, two agencies – Moody’s and Standard & Poor’s (S&P), founded in 1909 and 1916, respectively – were the dominant rating agencies of U.S. corporate debt. Other rating agencies, such as Duff & Phelps, which entered in the early 1970s, were considerably smaller. Fitch Ratings, although as old as the main agencies (it was founded in 1913 and has rated bonds on the AAA to D scale since 1924), was historically a much smaller agency.<sup>9</sup> Starting in 1989, and especially since its acquisition by a French investor in 1997, Fitch has invested in growing its market share to become an alternative to S&P and Moody’s, growing to more or less size-parity through both organic growth and acquisitions. Acquisitions include

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<sup>8</sup> This is a speculative argument. Revenues and fees are not public information. Even if there is no price premium ever paid for favorable ratings, amounts of future business could quite possibly be related to current ratings.

<sup>9</sup> See Cantor and Packer (1994) for evidence that Fitch still had a very low market share of the market for rating corporate securities in the early 1990-ies.

IBCA (British) in 1997, Duff & Phelps Credit Rating (American), and Thomson Bankwatch (Canadian) in 2000. Fitch's growth has varied considerably across industries, which is important to the construction of our empirical tests. Over the decade starting in the mid-1990s that we study, Fitch's share of corporate bond ratings issued has increased from around 10% to approximately one third of the market.<sup>10</sup>

We test theories of the impact of competition on reputation building, using the growth of Fitch's market share as the measure of competition faced by other rating firms. In our tests, we exploit the fact that entry varied across industries.<sup>11</sup> We rely on three different types of evidence. First, ratings issued by S&P and Moody's increased (moved closer to the top AAA rating) as competition increased.<sup>12</sup> Second, the correlation between bond yields and ratings fell. Third, we find evidence from equity price responses to firm downgrades. As competition increased, downgrade returns became larger in absolute value, consistent with a lower bar being set for the rating categories. We observe an even greater price decline in response to competitive entry if the rating downgrade takes the issuer from the investment grade category to the speculative grade one.

While we present three distinct pieces of evidence that suggest competition led to poor ratings quality, it is possible that our results are misleading due to some omitted variable. For instance, could Fitch's rate of entry in a particular industry be correlated with future changes in ratings levels, e.g., due to industry performance changes (beyond what's captured by firm controls)? One possibility is that a period of less-friendly credit ratings coincides with demand for alternative ratings by issuers who prefer not to see their ratings decline. We believe that this

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<sup>10</sup> This is the average across market shares at the 2-digit NAICS industry level. The same pattern is true for the median market share across these industries as well. In 2005-2006, Fitch's market share was highest in real estate, retail trade, utilities, construction and finance. Fitch remained relatively less represented in agriculture, educational services and transport. The largest gains (increase from 1997-98 to 2005-2006) were in public administration, real estate and construction, waste management and related, and retail trade.

<sup>11</sup> It's important for our empirical approach that industry-level market shares are informative about the competitive pressure across industries. Since rating agencies rely on extensive industry-specific knowledge to evaluate firm credit worthiness, it seems likely that the competitive situation varies across industries. Indirect evidence for this can be seen in the large number of industry-level reports published by the rating agencies.

<sup>12</sup> Blume, Lim, Mackinlay (1998) document a trend toward tougher ratings standards in the period preceding our sample.

is unlikely to explain our findings, however, since we find that Fitch's entry is correlated with friendlier ratings. Also, this explanation does not seem to explain why ratings become less correlated with bond yields as competition increases. We also test explicitly several theories of entry which might interfere with our interpretations. We find that Fitch's market share is not tied to increases or decreases in demand for debt in an industry, the amount of ratings issued in an industry, or industry profitability. Collectively, these additional tests lend greater credibility to our findings.

There are several implications of our findings. First, the fact that ratings quality seems to decrease with competition provides support for the standard economic theories of reputation (e.g., Klein and Leffler (1983)). Obviously, these findings do not necessarily indicate misbehavior by rating agencies, only that the equilibrium in the ratings industry relies on rents to reward reputation-building activities which are costly in the short run, and that the absence of such rents reduces the amount of reputation-building. Second, encouraging competition may reduce monopolistic (or in the case of ratings, oligopolistic) rents, but is not likely to improve quality. For policy makers, the benefits and costs of competition must be carefully compared.

There are several caveats and limitations to our findings. First, we only consider corporate ratings, not ratings of CDOs, mortgage-backed securities or other structured products.<sup>13</sup> Second, our findings have limited implications for the efficacy of reputation mechanisms in other imperfectly-competitive settings, since the ratings industry is a particularly special one. Third, we disregard many potential important aspects of reputation, such as how the reputational mechanism varies over firms' life-cycles (see Diamond (1989)) and how entrants appear in the industry (Mailath and Samuelson (2001)).

The rest of the paper is organized as follows. In Section 2, we discuss credit ratings and the underlying industry in more detail. In Section 3, we present the predictions of various

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<sup>13</sup> Doherty, Kartasheva, and Phillips (2008) examine the effect of competitive entry among rating agencies of the insurance market. Contrary to our findings in the corporate arena, they find that the entry of S&P as a competitor to the incumbent monopolist of A.M. Best led to improved rating content.

theories, and the methodology used to test them. We present the data in Section 4 and results in Section 5. Concluding remarks can be found in Section 6.

## 2. Credit ratings: background and policy debate

Credit ratings range from AAA to D (see Table 1B for an overview of the ratings levels for the three main rating agencies and our numerical value assignments for our empirical work).<sup>14</sup> There are two main types of ratings. Bond ratings are provided for a vast majority of publicly-traded bonds in the United States (U.S.). Firm (or Issuer) ratings are produced by each of the three main agencies for all U.S. public firms that issue public debt. Ratings are typically shared freely by the rating agencies; whose revenues derive from charges to the firms whose credit quality is being assessed. Fees for bond ratings typically consist of a fixed fee per year coupled with a larger upfront fee which is charged when the bond issue is first rated at time of issuance.<sup>15</sup> Paying for firm ratings is voluntary, although raters will *only* consider non-public information provided by the firm itself if they receive payment from the corporate issuer (see Jorion et al. (2005) regarding raters' access to non-public information). Ratings generated at the request of the issuer are referred to as solicited ratings for which the aforementioned fees apply, whereas ratings assigned not at the issuer's request are called unsolicited. Rating agencies also provide various other types of ratings, such as short-term credit opinions and various industry-specific ratings.

Since 1975, the Securities and Exchange Commission (SEC) has limited competition in the market for credit ratings by designating only certain firms as "Nationally Recognized Statistical Rating Organizations" (NRSROs). This may induce concerns for maintaining a reputation of quality ratings beyond that induced by investors and their demand for high quality ratings. It may also make entry in the industry more difficult, since many investors will only consider ratings by an NRSRO when making investment decisions.<sup>16</sup> Some argue, such as

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<sup>14</sup> Sylla (2002) presents an excellent history of the ratings industry in the US.

<sup>15</sup> Fees vary with the face value of a bond issue, but usually in a non-linear way (i.e., they are capped). Also, active issuers may receive quantity discounts. In February, 2008, S&P shared information about their rating fee structure, including that corporate issuers (including industrial and financial service companies) pay "up to 4.25 basis points for most transactions" and that the minimum fee is \$67,500. Also, "S&P will consider alternative fee arrangements for volume issuers and other entities that want multi-year ratings services agreements" (Standard and Poor's 2008).

<sup>16</sup> See Boot, Milbourn and Schmeits (2006) for both a discussion and model of such investor restrictions to hold only investment grade debt securities.



SEC Commissioner Paul S. Atkins, that “the unintended consequence of the SEC's approach to credit rating agencies was to limit competition and information flowing to investors. The legislative history reflects a genuine concern that the SEC facilitated the creation of – and perpetuated – an oligopoly in the credit rating business. Indeed, today, three NRSRO-designated firms have more than 90 percent of the market share.”<sup>17</sup>

To what extent might rating agencies diminish either the quality or informativeness of their ratings to garner further market share? It has in fact been alleged that Brian Clarkson, upon being named President of Moody's in August 2007, “set out to make [the firm] more client-friendly and focused on market share”.<sup>18</sup> In recent US Congressional hearings, practitioners have spoken out against the rating agencies and the SEC's process for recognizing these agencies, of which there are currently only five with the esteemed designation of NRSROs. The President of the Investment Company Institute, Paul Schott Stevens, said during Senate testimony that “unfortunately, the current designation process does not promote – but, in fact, creates a barrier to – competition.”<sup>19</sup>

### **3. Theory, hypotheses and methodology**

The key empirical challenge in this study is to define rating quality in a theoretically appealing and empirically relevant manner. The SEC (2003) uses the phrase that they want to promote a market environment resulting in “credible and reliable ratings”. In our attempt to understand the impact of increased competition in this industry, we use several complementary approaches and three main datasets to evaluate rating quality. Our methods are based on the idea that high quality ratings should be accurate, informative and honest. Hence, we assume that lower quality ratings – that is, ratings more influenced by issuer preferences – will be more favorable to issuers (i.e., higher) and less informative about credit quality. There are multiple implications of this statement. First, lower quality ratings will be on

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<sup>17</sup> See “Speech by SEC Commissioner: Remarks to the Institute of International Bankers”, by SEC Commissioner Paul S. Atkins, March 3, 2008. Link to full speech is here: (<http://www.sec.gov/news/speech/2008/spch030308psa.htm>)

<sup>18</sup> See “Rating Game – As Housing Boomed, Moody's Opened Up”, *Wall Street Journal*, April 11, 2008, page A1.

<sup>19</sup> See <http://www.financial-planning.com/asset/article/527499/fund-industry-group-calls-more-credit.html>

average better ratings, that is, ratings closer to the AAA end of the spectrum since this must be the universal desire of the issuers as the subject of the ratings.

Second, we also use information in equity prices. A firm's stock price tends to fall on announcement of a rating downgrade (i.e., the announcement that a firm's or bond's rating has been lowered by a rating agency).<sup>20</sup> If ratings standards deteriorate as competition increases, downgrades should be worse news in the wake of increased competition, since the downgraded security (or firm) has failed to pass an even lower quality bar than what was in place originally.<sup>21</sup> The return to ratings downgrades should then be more negative (i.e., larger in an absolute sense) in this issuer-friendlier environment. Conversely, if ratings standards improve as competition increases, ratings downgrades should result in smaller equity price drops because a downgrade under stricter quality standards suggests less negative news. We test these predictions by examining if the negative equity returns around downgrades are smaller or larger when there is more competition. It is worth pointing out that, unlike the first prediction, this cannot be a permanent effect. In the short run, downgrades are worse news, but in the long run, the ratings distribution has adjusted, and downgrade return should revert in magnitude.<sup>22</sup>

Third, lower quality ratings mean that ratings will reflect things other than expected repayment, and thereby rating levels will likely be less correlated with bond yields. Testing the informativeness of ratings is slightly more challenging. Direct testing using actual performance, such as the observed default rates and repayment histories for various rating categories, is impractical. Such direct testing would require a very long data horizon (since many ratings are issued for securities with long maturities for which ultimate payment performance is unobserved for a very long time). Second, for most rating categories, default is very unusual. This doesn't mean that individual ratings are not distinct, only that the

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<sup>20</sup> See, for example, Jorion et al (2005).

<sup>21</sup> This effect is likely to be temporary, since in the longer run, all ratings will be set consistently with these lower (less informative) standards.

<sup>22</sup> It may seem reasonable that Fitch can only exert pressure on the ratings of S&P and Moody's if Fitch itself on average offers friendly ratings. Fitch ratings are not included in our samples, but in test (not reported) comparing Fitch bond ratings to those issued by the other two rating agencies, Fitch's ratings are 0.2 steps higher (controlling for bond fixed effects), consistent with competition through ratings levels.

distinction is difficult to identify using actual defaults. Instead of actual payment performance, we use market prices of debt to assess the informativeness of ratings. We examine the correlation of ratings with bond yields, conditional on various controls known to correlate with yields. That is, we ask if ratings contain information about bond values beyond easily observable characteristics such as bond covenants and firm characteristics. In particular, we test if competition reduces or increases the informativeness of ratings, as measured by the conditional correlation with yields.

If ratings in practice are fraught with imperfections (they become of lower quality and informativeness), then this is most likely the case because the subject of these ratings (the issuing firm) has a strict preference for more favorable ratings. However, there should be some cross-sectional variation among firms with regards to their preferences for better ratings. In particular, more heavily indebted firms are likely to care more about ratings than less indebted firms (this is born out in survey data presented in Graham and Harvey (2001)). We therefore exploit cross-firm variation in the importance of ratings to issuers, and predict that any effect of competition would be stronger for firms with higher leverage.

Our tests rely on the use of Fitch's market share as a measure of competition. We calculate this based on the number of bond ratings issued. This is not a perfect measure of market share, and revenue share would probably be preferable, but is not readily available. We believe that Fitch's market share of bond ratings is indicative of the competitive threat to S&P and Moody's in a segment of the market for ratings. One advantage of using individual bond ratings is that it affords us a very large dataset. As described in the next section, our sample covering the early 1990s through 2007 yields a total of approximately 1.1 million ratings.

#### **4. Data**

Our tests require drawing data from a number of sources. We also collect data on credit ratings for individual bonds (the issues) as well as for firms (the issuers), on firm characteristics and accounting numbers, on equity returns around rating downgrade events, and on bond yields for rated securities.

Data on bond ratings and market shares are drawn from the Mergent Fixed Income Securities Database (FISD). This database provides both issue- and issuer-specific data. We use data on ratings by S&P, Moody's and Fitch of individual issues (bonds) to estimate the market share of Fitch in each industry-year cell. The total number of bond ratings used to calculate market shares is approximately 1.1 million. Each bond rating is matched to an industry using the issuer's Cusip. There are more ratings around the year 2000 than in other years, but no year has fewer than 30,000 ratings. We define Fitch's market share as the fraction of all bond ratings in a year-industry cell performed by Fitch, where industries refer to the 2-digit North American Industry Classification System (NAICS) industries and our sample years run from 1995 to 2006 (although many of our tests will not use the first few years of data). Figure 1 presents a moving average of monthly market shares for Fitch from 1998 to 2006.<sup>23</sup> Fitch's market share increases especially fast in 2000, coinciding with two acquisitions.<sup>24</sup>

Firm ratings and accounting data are collected from the Compustat Industrial and Operating Segments databases. Compustat also contains S&P issuer credit ratings, defined as "a current opinion of an issuer's overall creditworthiness, apart from its ability to repay individual obligations. This opinion focuses on the obligor's capacity and willingness to meet its long-term financial commitments (those with maturities of more than one year) as they come due". We examine these ratings as well as individual bond ratings by Moody's and S&P. An important caveat to the use of these Compustat long-term debt ratings is the fact that they are updated only annually. In other tests below that require more precise calendar information on ratings changes, we rely on another database.

To identify bond yields, we use bond transaction data from the Mergent FISD database. This dataset covers all bond acquisitions and disposals (sales, redemptions) since 1995 by insurance companies. We exclude bonds denominated in foreign currencies, as well as any bonds that are callable, puttable, convertible, substitutable or exchangeable. We also exclude

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<sup>23</sup> In tests, we use the total market share for each industry-year. This figure presents moving averages of total monthly market share across industries in order to provide a sense of the time path of Fitch's entry.

<sup>24</sup> Potentially, market share increases due to organic growth and acquisitions have different competitive impact, and including data from 2000 may make our results less representative. We have rerun our ratings levels regressions (Tables 3, 4, 6) using only post-2000 data, or all years except 2000, with very similar results.

US issues by foreign issuers (i.e., Yankee bonds). We drop defaulted bond issues, bonds denominated in foreign currency, and bonds with refund protection. We drop variable coupon bonds (because their yields to maturity are harder to calculate). We also require several control variables (such as issuer industry) to be available, and drop bond trades with very high or very low sales prices to avoid data errors (this constraint does not affect our results). Most of these restrictions do not reduce the sample size much.

We match each bond transaction to the most recent rating of the bond by Moody's or S&P, and throw out any bonds with no ratings in the month preceding the transaction. If there is more than one rating on the same date, we use the median of the most recent ratings. The remaining sample of bond transactions consists of a little more than one hundred thousand observations. For these we estimate the yield-to-maturity numerically to match the observed sales price. The average yield to maturity is 6.5%. We also calculate yield spreads by subtracting the yield on government bonds of similar maturity. Government yield data is from Federal Reserve H15 reports. For each bond in this sample, we also try to identify the initial issue yield, and match that to an early rating.

We also collect data on firm upgrades and downgrades, which are changes in firm (issuer) ratings. The source for these data is Standard & Poor's Ratings History. We hand-match this sample to CRSP to generate a company-level Permno. Through this process, we obtain a total of 1,585 issuing firm credit rating downgrades with some matching stock return data for 543 different firms (our number of observations in some regressions is somewhat reduced due to limited availability of independent variables). Of these downgrades, 221 are downward movements from investment grade (BBB- and higher) to junk grade (BB- and lower), which has an important practical consequence for investor participation.

An overview of the most important variables is presented in Table 1A. The number of observations for Fitch's market share refers to the number of industry-year cells.

## **5. Empirical results**

This section presents our evidence from the various tests of rating quality and how it is seemingly affected by changes in the competitive landscape of rating agencies.

### *5.1. Bond and firm credit rating levels*

The first test of rating quality and how this is affected by increased competition is for the level of firm credit ratings. We regress firm ratings on Fitch's market share. Results are presented in Table 2. In column one, no controls are included. In this sample, there is a significant positive correlation between competition and credit ratings, suggesting that more competition pushes ratings toward the higher end of the rating spectrum (i.e. toward AAA). This pattern is clearly illustrated in Figure 2 which plots the frequency of each rating for industry-years with high and low values for Fitch's market share. As the graph shows, all investment grade ratings (i.e. BBB- and above) are more common under high competition, and all junk bond ratings (i.e. BB+ and below) are more common under low competition. In other words, the figure and the regression analysis offer complementary evidence that competition is correlated with higher ratings. The result that competition coincides with ratings that are friendly to raters, which we interpret as a suggestion of lower quality, is consistent with theories (along the lines of Klein and Leffler (1984)) that predict a negative effect of competition on product quality. The result in column one may be unreliable, however, since no controls are included.

In column two, we rectify this by including year and industry dummies. This pushes up the R-squared significantly, and reduces the coefficient and standard error on competition. In this specification, the coefficient on Fitch's market share remains positive and significant. The magnitude is modest but non-trivial. For a one standard deviation change in competition (0.142), average ratings are predicted to increase by 0.19. This corresponds to a one rating step upgrade (e.g., BBB+ to A-) of approximately one out of every five firms. In column three, we include firm fixed effects (which makes industry fixed effects redundant), which absorb most of the variation (firm ratings are fairly stable). The estimated effect of competition remains positive and significant, and the implied magnitude is slightly smaller (one in nine firms). So far, we have not controlled for any time-varying features of a firm. In column four, we also include further firm controls (see table notes) intended to capture time variation in firms' performance. The estimated coefficient on Fitch's market share is very similar to the one found in the previous specification.

Implicitly, the left hand side variable treats every step of the rating system as equal in the OLS specifications (see Table 1B for details of the numerical rating variable). There is no reason for this to be accurate, however. In column five, we run an ordered probit regression instead of OLS. This specification allows each cut-off to be estimated and so implicitly allows the effect of dependent variables to vary across different levels of ratings (although it may be less robust to certain econometric problems than OLS). The effect of competition remains positive and significant.

Overall, the firm rating results suggest that ratings become more favorable to issuers when competition increases, consistent with Klein and Leffler (1983) style theories that suggest disciplining effects of reputation are diminished as competition increases.

As a robustness test, we turn now to ratings of individual issues as opposed to the issuing firms. Such tests should provide further evidence of how increases in competition among rating agencies can affect the quality of ratings. In Table 3, we report the estimates of regressions of bond credit ratings on Fitch's market share. The number of observations is very large, since many firms issue very many bonds. We include a range of fixed effects in order to control for the effects of observables on bond ratings. In column one, we report a regression of ratings on Fitch's market share, controlling for both year and industry fixed effects, as well as the previous rating for the same bond. Together with our competition measure, the fixed effects capture 94% of the variation in ratings. The effect of competition is positive and significant, in line with the finding for firm level ratings. Observe that the estimated magnitude is about half of the firm-level effect (a one standard deviation increase in competition predicts that one in twenty one bonds will have a one step higher rating).

The large number of observations of bond ratings affords us even more fine-tuned controls. In column two, we control for bond duration non-parametrically by including fixed effects for time to maturity (measured in years) as well as the lagged rating (i.e., the previous rating by S&P or Moody's, whenever it occurred). The coefficient estimate is now slightly higher (one in sixteen bonds have a predicted rating increase). In column three, we include bond issue fixed effects, which make industry fixed effects redundant. In this specification, Fitch's market share is again positively and significantly related to ratings (this time, the

implied magnitude is a one step upgrade for one in fourteen bonds, and the significance level is 5%). In column four, we exclude financial firms (NAICS 52) from the sample in case bonds in this industry are fundamentally different from non-financial firms' bonds. The sample size is cut in half, but we get results that are similar to the full sample regressions. The estimated coefficient is larger for the non-financial sample, corresponding to an upgrade of one in eight firms for a one standard deviation increase in Fitch's market share. In all our bond ratings regressions, ratings increase with competition, consistent with the theories predicting a negative effect of competition on quality and in agreement with the firm-level results.

### *5.2. Firm credit rating levels: interaction tests*

We next consider cross-sectional variation in the impact of competition on firm ratings. The effect of competition should be felt more acutely for those firms that are likely to care more about their ratings. We use firm indebtedness to identify firms with a greater concern for ratings. In Table 4, we interact Fitch's market share with four measure of indebtedness: leverage (debt over assets), long-term leverage (long-term debt over assets), a high leverage dummy (leverage is above the median in the firm's industry) and debt divided by EBITDA. These specifications allow us to include industry-year interaction fixed effects (i.e., approximately 400 dummies), thereby reducing any concern about omitted variables that are correlated with Fitch's market share and vary within industries and years. Without exception, the interactions of competition and debt are negative and highly significant. This suggests that the effect of competition is disproportionately felt for firms which are likely to care more about their ratings because they rely more heavily on debt financing. This is consistent with the argument that competition makes ratings more responsive to firm preferences.

### *5.3. Bond yields and ratings*

With the empirical results related to rating levels in hand, we turn now tests of how informative bond ratings are and whether this informativeness changes in response to competition. In particular, we examine the conditional correlation of yields and ratings declines when competition increases. We test this by including Fitch's market share times a bond's credit rating in a regression of bond yields on bond characteristics. Results are reported in Table 5. Bond trades occur at different times, and interest rates are likely to be an important source of time series variation in yields, so we include fixed effects for each date (specifically, each month-year pair). In column one, we include a bond's most recent credit rating and Fitch's



market share in that industry year as controls. We also include fixed effects for industries as well as controls for bond characteristics (time to maturity and size of bond issue). The coefficient on credit ratings is negative and significant, confirming that bonds with better credit ratings trade at lower yields. The coefficient on the interaction of credit rating and Fitch's market share is positive and significant, implying that the correlation of credit ratings and bond yields is lower when competition is stronger. The magnitude of this effect is economically large. A one standard deviation increase in Fitch's market share reduces the coefficient on credit ratings by approximately a fifth of the implied value when Fitch has zero market share. This is consistent with the view that competition *reduces* the information content of ratings.

In column two, we also include year-industry fixed effects (this precludes identifying the coefficient on Fitch's market share). The results agree with the previous regression, and if anything imply a slightly stronger impact of competition on the correlation of yields and ratings. Increasing Fitch's market share by one standard deviation reduces the coefficient on credit ratings by about a fifth of the implied coefficient with zero Fitch market share. Alternatively, the coefficient on credit ratings is about a third lower when moving from the 25<sup>th</sup> to the 75<sup>th</sup> percentile of the competition measure.

One concern with the specification in column (2) is that we have not perfectly controlled for yield curve changes. We therefore deduct the yield on the closest maturity Treasury bond for each observation, and using the resulting yield spread as dependent variable. This yield spread is 138 basis point on average, with a standard deviation of 134 basis points. The regression produces results that are very similar to the previous specification, both in terms of magnitude and significance (the sample is slightly smaller when no match can be found, e.g. because the bond's maturity is too long).<sup>25</sup>

These tests are base on trades of bonds, where the same bond can appear multiple times (approximately five times on average). One concern is that multiple trades capture the same information. We include bond controls (and in unreported regressions, have included bond

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<sup>25</sup> We have also included the interaction of date (i.e. month-year) fixed effects and the natural log of time to maturity to absorb any variation in how credit spreads vary with bond age (not reported). Also, we have included controls based on estimated bond durations instead of maturities (not reported). These variations have a very small impact compared to the reported regressions.

issue fixed effects).<sup>26</sup> However, a more direct way of addressing this is to look at the price at issue. The Mergent/FISD database contains yield spreads at issue and we can match these to early ratings using the same process as for the secondary market trades. In column four, we use these yield spreads at issue. As with data on secondary market trades, the implied correlation between credit ratings and bond yields is weaker when Fitch's market share is higher. The magnitude is similar to, but slightly lower than, that found in the trade data.

An alternative way of assessing the correlation between yields and ratings is to run first stage regressions with all controls except ratings, and then regress residuals from the first stage on bond ratings. The second stage can be done separately for subsamples split by Fitch market share. Such specifications, containing multiple stages, are harder to assess statistically than our interaction tests. However, they give very similar results. For example, we have run a two step version of the regressions in column one (not reported), dropping ratings and the Fitch market share-rating interaction from the first stage. We then regressed yield residuals on ratings in a second stage, having split the sample in half by competition. For the high market share sample, the R-squared is 0.086, and for the low market share sample, 0.141.<sup>27</sup> In other words, the explanatory power of ratings for yields is higher when Fitch's market share is low. Again, ratings are more informative when competition is low.

The results in Table 5 show that the correlation of ratings and yields declines as competition increases. In other words, bond yields (and spreads) are less related to credit ratings when Fitch has a high market share. The implication is that credit ratings are less informative for yields when competition is stronger, or contain less yield-relevant information. Overall, this supports the theories that predict lower quality (less informative) ratings when there is more competition.

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<sup>26</sup> We have tried clustering errors by bond issue, and consistently get much higher significance than reported in the tables (where errors are clustered by industry-year combinations).

<sup>27</sup> The estimated coefficients on ratings are -0.094 and -0.215, respectively, and both highly significant.

#### 5.4. Rating downgrade announcement returns

In the final set of tests, we explore the information content of ratings by examining equity price reactions in response to rating downgrades. We follow the general methodology of Holthausen and Leftwich (1986).<sup>28</sup>

A quality reduction in ratings could have two contradictory effects. First, lower quality ratings should correspond to reduced creditor quality in any particular ratings category. One can then interpret this as the rating agency “lowering the bar”, making a rating downgrade now convey worse news since the firm is falling below an even lower quality threshold. This should tend to make downgrade returns more negative (i.e. larger in an absolute sense) in response to an announced rating downgrade.

There is also a second, long-term effect of such quality shifts. This relates to the reduced overall informational content once the population of ratings has come to reflect these factors other than credit quality. Lower quality ratings reflect factors other than expected repayment on the issue and that such ratings will be less correlated with the creditworthiness of borrowers. This will likely tend to reduce the amount of information released with downgrades, hence moving downgrade returns toward zero. Since this effect requires the population of ratings to be moved to a new equilibrium, whereas in practice revising ratings is most likely a slow and gradual process, this effect is unlikely to be visible quickly. It is therefore likely to be more challenging to detect econometrically.

We rely on this distinction in timing and focus on the first effect, which implies larger (i.e., even more negative) downgrade returns as competition increases. If our competition measure correlates with worse quality of ratings, returns should be negatively correlated with competition. We present results for this type of test in Table 6.

The dependent variable in Table 6 is the equity return during an event window around a firm downgrade, net of the market return during the same period.<sup>29</sup> We use daily event

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<sup>28</sup> See Jorion et al (2005) for a recent example. There is also a literature looking at bond price reactions to downgrades, including Weinstein (1977), Wakeman (1978), Katz (1974), Grier and Katz (1976) and Ingram, Brooks and Copeland (1983) (with mixed findings). Hand, Holthausen and Leftwich (1992) find excess returns of around -0.80% for the day of and day after a downgrade announcement.

windows of [-1,1], [-2,2] and [-10,1] (the latter window is examined in case there is pre-event return drift). In columns one to three, we control only for industry and year fixed effects. The effect of competition on returns is negative for all event windows, and the two longer event windows present significant coefficients. The implied magnitude is also large. Based on column two, an increase of Fitch's market share is implied to reduce average event returns by about half of the mean, or 15% of a standard deviation.<sup>30</sup>

Event returns are likely to vary by firm, and reflect features such as the variability of its share price. In column four, we control for firm volatility, along with its square and cube to capture any nonlinearities. Fitch's market share has a negative and significant effect on five day event window returns, with a magnitude essentially unchanged from column two. The distinction between junk and investment grade is often considered particularly important. In our sample, the mean equity return in a five day event window is 3.7% for such downgrades, slightly higher than the average 2.7% return for all downgrades. By focusing on these larger events, we hope to more clearly identify the effect of competition, although it will reduce our sample size. In columns five and six, we focus exclusively on downgrades from investment grade to junk status, leaving us with 182 observations. With or without controls for volatility, in column five and six, respectively, we find an effect of competition about three times as large for downgrades to junk as for the full sample. The effect is more significant than for the full sample. A one standard deviation increase in Fitch's market share implies reduced average event returns by slightly more than the mean, or 40% of a standard deviation.

The results in Table 6 suggest that competition has made ratings more lenient and lowered the quality bar for downgrades, making announced downgrades worse news and equity returns around such downgrades more negative.

### *5.5. Alternative explanations*

The empirical strategy of this paper relies on industry-level variation in the extent to which Fitch competes with the incumbent rating agencies. We have suggested that the rate of

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<sup>29</sup> We have run similar tests with equity returns around bond downgrades. Although mean returns for these are different, the result for the effect of competition is very similar.

<sup>30</sup> We have also clustered standard errors by firm instead of by industry-year, yielding similar results (but with somewhat higher significance).

entry is largely driven by factors that do not relate specifically to industry ratings quality, and therefore we needn't worry about omitted variables and can treat Fitch's market share as exogenous. Factors that might determine the rate of entry are the speed with which suitable staff can be found and the accumulation of industry expertise. The key assumption is that entry is not driven by factors such as the growth in industry demand for credit or for credit ratings, or by variables that might be correlated with ratings quality through channels other than competition. In particular, we are concerned that Fitch tends to get a relatively higher market share when an industry's credit demand is growing rapidly (and hence the industry's demand for ratings). In this case, we might simultaneously see improved ratings (if credit demand coincides with good industry performance) and lower risk premia (which might manifest itself as a lower correlation between prices and ratings).<sup>31</sup> Credit demand growth may not be directly observable, but we can imagine several factors that might correlate with it, such as the amount of credit outstanding in an industry, the average growth rate, the number of ratings issued, and industry profitability.

We attempt to examine alternative explanations for Fitch's relative rate of entry in different industries explicitly. To do this, we examine the extent to which Fitch's market share is correlated with the contemporaneous value, as well as various lead and lag values, of the number of ratings in an industry, the level and changes of industry debt outstanding, and industry profitability (weighted average and median). For these tests, we regress industry-year observation of one of the variables on Fitch's market share, controlling for industry and time fixed effects. We repeat this for Fitch's market share timed at different lags and leads. Results are presented in Table 7. There seems to be no correlation between the various measures of demand for credit and Fitch's market share (one coefficient out of twenty five is estimated to be significant at the 10% level, which is slightly fewer than the expected number from a random sample). We conclude that Fitch's relative market share, controlling for industry and year fixed effects, is unlikely to reflect systematic demand factors and more likely to be driven by idiosyncratic factors internal to the firm or related to labor markets.

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<sup>31</sup> This alternative story is more difficult to reconcile with our findings on equity announcement effects, but we wish nevertheless to consider it carefully.

Another alternative explanation for ratings that improve when there are more raters present in a market is ratings shopping. “Ratings shopping” refers to a situation when an issuing firm asks raters for opinions on the rating they might receive for a particular issue, and then choose to get ratings only from the rater(s) promising the most favorable ratings. In the absence of any biases, but as long as rating agencies disagree, the average rating might be higher when there are more raters to choose from. There are three reasons why we believe this to be unlikely. First, ratings shopping would not explain our findings in the issuer-level analysis (cf. Table 2) since these ratings all come from the single rating agency of S&P. Second, there is little evidence for ratings shopping among US bond issuers (see Cantor and Packer (1997) and Jewell and Livingston (1999)). Finally, if ratings shopping were a key factor in ratings demand, we would expect to see Fitch’s market share increase when industry performance was poor or when ratings were declining. Neither of these patterns is observed empirically (cf. Tables 2, 3 and 7).

## **6. Conclusions**

Credit ratings are a key aspect of the financial system. The quality of ratings is certainly sustained in part by the reputational concerns of rating agencies, whose paying customers have no inherent interest in the quality of ratings. Competition in this industry has been increasing and there have been calls for yet more competition. Will this reduce quality, as can be predicted by an argument along the lines of Klein and Leffler (1983) or improve it, as perhaps predicted by Hörner (2002)? We test these conflicting predictions in the ratings industry using the entry of Fitch Ratings as an experiment in the amount of competition faced by the incumbent rating agencies of S&P and Moody’s.<sup>32</sup>

We find three pieces of evidence, all more or less consistent with a reduction in credit rating quality as Fitch increased its market presence. First, competition is associated with friendlier ratings (i.e., they are closer to AAA). Second, ratings and bond yields have become

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<sup>32</sup> The system of third party ratings is based on considerable investment by rating agencies in a reputation for honesty and precision. These investments are only likely to occur if the rewards are commensurate. The current system relies on the existence of rents outweighing the short-term interest of individual issuers. Our study confirms this, but implies no criticism of individual firms. The expectation that rating agencies should provide a public good for free is unrealistic. If they are to fulfill their function, rents may be necessary.

less correlated (conditional on public information about bonds and issuers). Third, at least in the short run, equity prices react more to downgrades as competition increases, consistent with a lowering of the bar for ratings categories. This is especially clear for downgrades from investment grade to junk status.

The economic magnitudes we find are moderate but nontrivial. Conservatively, we find that a rise in competition corresponding to a one standard deviation increase in Fitch's market share is predicted to increase the average firm and bond rating by 5-10% of a rating step (and increase it significantly more for more highly-levered firms), to reduce the conditional correlation between ratings and bond yields by about a sixth compared to the case when Fitch has no market share, and increase the negative equity price responses to downgrades by a quarter or more.

These results have potential policy implications. For regulators, it is worth considering that increasing competition in the ratings industry involves the risk of impairing the reputational mechanism that underlies the provision of good quality ratings. There may obviously be benefits of competition in other areas (e.g., reducing rents may be a policy goal in and of itself). Nevertheless, calls for more competition, such as by the U.S. Department of Justice (1998), may deserve a caveat. For bond markets, it is clear that relying on third party ratings paid for by issuers is not a system without risks. Our empirical findings suggest that the system will work better when competition is not too severe. These results about the level of competition and the efficiency of reputational mechanisms offer support for models of the Klein and Leffler (1983) variety. In other words, competition reduces future rents and increases the short terms gains to cheating, and hence makes the reputational equilibrium harder to sustain. Obviously, these implications may not apply to other markets and in other settings.

## References

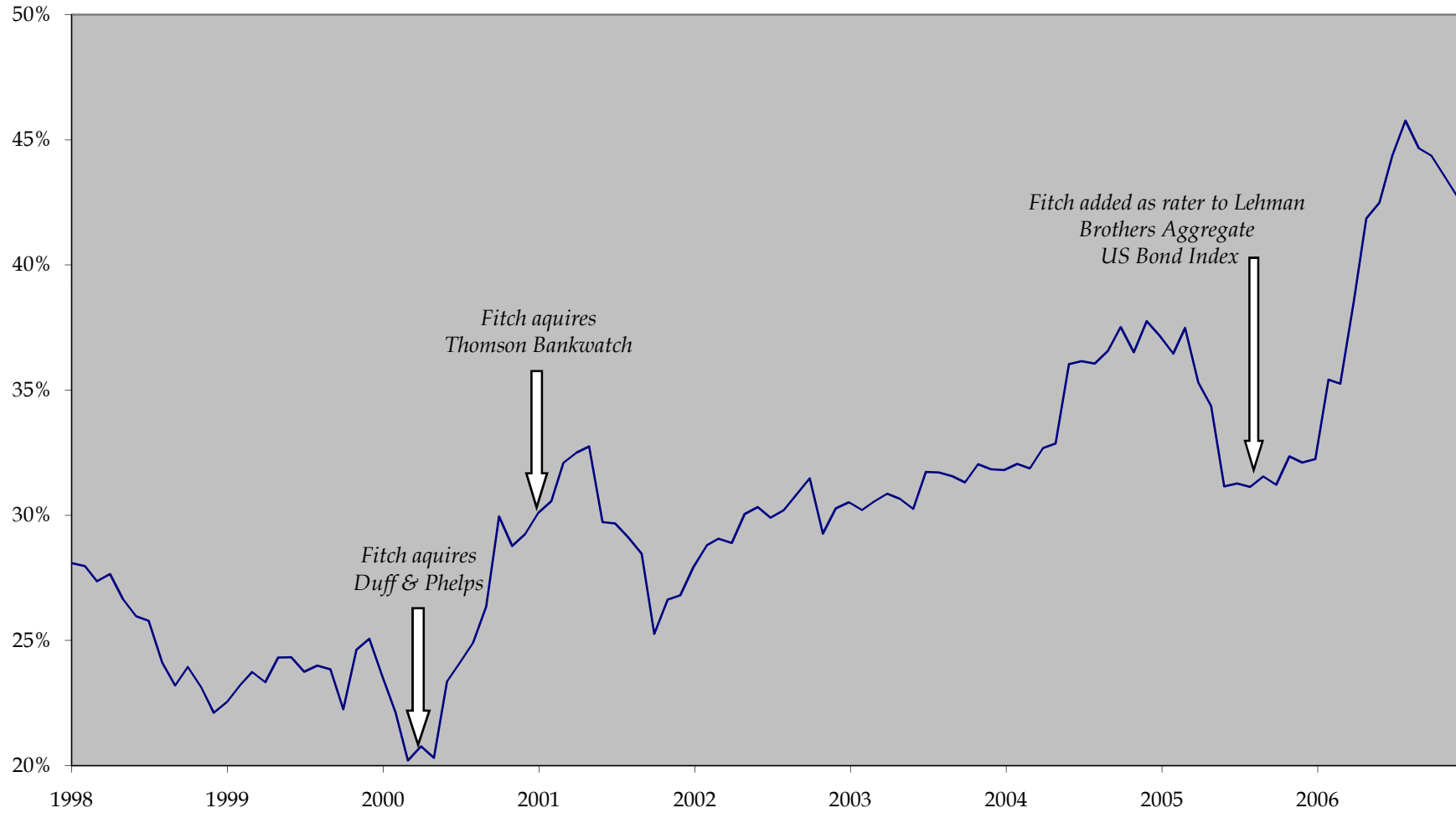
- Bar-Isaac, Heski, 2005, "Imperfect Competition and Reputational Commitment", Economic Letters, 89(2), November, 167-173.
- Bar-Isaac, Heski, and Steven Tadelis, 2008, "Who wants a reputation", Foundation and Trends in Microeconomics, forthcoming.
- Blume, Marshall E., Felix Lim, and A. Craig Mackinlay, 1998, "The Declining Credit Quality of U.S. Corporate Debt: Myth or Reality?", Journal of Finance, 1389-1413.
- Boot, Arnoud W.A., Todd Milbourn, and Anjolein Schmeits, 2006, "Credit Ratings as Coordination Mechanisms", Review of Financial Studies 19-1, 81-118, Spring 2006.
- Cabral, L. and A. Hortaçsu (2006), "The dynamics of seller reputation: Theory and evidence from eBay", Mimeo, NYU
- Campbell, John Y. and Glen B. Taksler, Equity, 2003, "Volatility and Corporate Bond Yields", The Journal of Finance, Vol. 58, No. 6. (Dec.), pp. 2321-2349.
- Cantor, Richard, and Frank Packer, 1997, "Differences of Opinion and Selection Bias in the Credit Rating Industry", Journal of Banking and Finance, 21, pp. 1395-1417.
- Chevalier, J. and G. Ellison, 1999, "Career concerns of mutual fund managers", Quarterly Journal of Economics, 114, 389-432.
- Hong, H. and J. D. Kubik, 2003, 'Analyzing the analysts: Career concerns and biased earnings forecasts'. Journal of Finance 58, 313-351.
- Hong, Harrison and Marcin Kacperczyk, 2008, "Competition and bias", working paper.
- Hubbard, T. N., 2002, "How do buyers motivate experts? Reputational incentives in an auto repair market", Journal of Law and Economics, 45, 437-468.
- Jin, G. Z. and P. Leslie, 2003, "The effect of information on product quality: Evidence from restaurants hygiene grade cards", Quarterly Journal of Economics 118, 409-451.
- Jin, G. Z. and P. Leslie, 2008, "Reputational incentives for restaurant hygiene", American Economic Journal (Microeconomics), forthcoming.
- Cantor, Richard and Frank Packer, 1994, "The Credit rating Industry", Federal Reserve of New York Quarterly Review, 19(2), Summer-Fall, 1-26
- Cooper, Russell and Thomas W. Ross, 1984, "Prices, Product Qualities and Asymmetric Information: The Competitive Case." Review of Economic Studies, April, 51(2), pp. 197-207.
- Department of Justice, 1998, Comments of the United States Department of Justice in the Matter of: File No. S7-33-97 Proposed Amendments to Rule 15c3-1 under the Securities Exchange Act of 1934, (March 6).
- Diamond, Douglas W, 1989, "Reputation Acquisition in Debt Markets", The Journal of Political Economy, Vol. 97, No. 4 (Aug), pp. 828-862



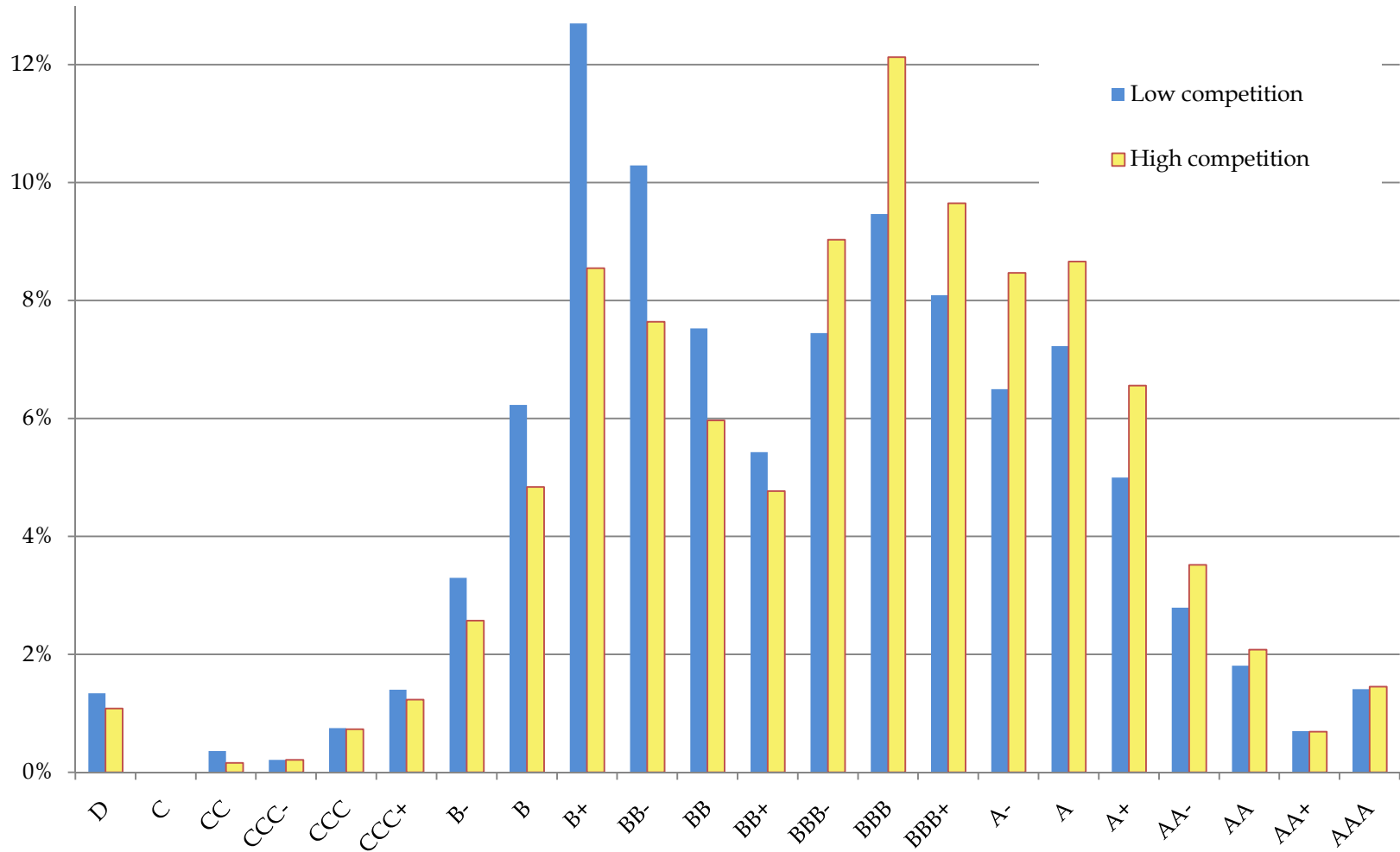
- Doherty, Neil A., Anastasia Kartasheva, and Richard D. Phillips, 2008, "Does Competition Improve Ratings?", working paper.
- Graham, John and C Harvey, 2001, "The theory and practice of corporate finance: evidence from the field", Journal of Financial Economics, 60(2-3), 187-243.
- Grier, Paul and Steven Katz, 1976, "The Differential Effects of Bond Rating Changes Among Industrial and Public Utility Bonds by Maturity", The Journal of Business, Vol. 49, No. 2. (April), 226-239.
- Hand, John R. M., Robert W. Holthausen and Richard W. Leftwich, 1992, "The Effect of Bond Rating Agency Announcements on Bond and Stock Prices", The Journal of Finance, Vol. 47, No. 2 (June), pp. 733-752.
- Holthausen, Robert W. and Richard W. Leftwich, 1986, "The Effect of Bond Rating Changes on Common Stock Prices", The Journal of Financial Economics, Vol. 17 (September), pp. 57-89.
- Ingram, Robert W., Leroy D. Brooks; Ronald M. Copeland, 1983, "The Information Content of Municipal Bond Rating Changes: A Note", The Journal of Finance, Vol. 38, No. 3. (June), pp. 997-1003.
- Jewell, Jeff and Miles Livingstone, 1999, "A Comparison of Bond Ratings from Moody's S&P and Fitch IBCA", Financial Markets, Institutions & Instrument, Volume 8 Issue 4, Pages 1 - 45.
- Jorion, Phillippe, Zhu Liu and Charles Shi, 2005, "Informational effects of regulation FD: evidence from rating agencies", Journal of Financial Economics, Volume 76, Issue 2, May 2005, Pages 309-330.
- Katz, Steven, 1974, "The Price and Adjustment Process of Bonds to Rating Reclassifications: A Test of Bond Market Efficiency", The Journal of Finance, Vol. 29, No. 2, Papers and Proceedings, (May), pp. 551-559.
- Kisgen, Darren J., 2006, "Credit Ratings and Capital Structure", The Journal of Finance, Vol. 61, No. 3, (May), pp. 1035-1072.
- Klein, Benjamin and Keith B. Leffler, 1984, "The Role of Market Forces in Assuring Contractual Performance", The Journal of Political Economy, Vol. 89, No. 4. (August), pp. 615-641.
- Mailath, George J., and Larry Samuelson, 2001, "Who Wants a Good Reputation?", The Review of Economic Studies, Vol. 68, No. 2, April, pp. 415-441
- Securities and Exchange Commission, 2003, Report on the Role and Function of Credit Rating Agencies in the Operation of the Securities Markets As Required by Section 702(b) of the Sarbanes-Oxley Act of 2002.
- Shapiro, Carl., 1983, "Premiums for High Quality Products as Returns to Reputations." Quarterly Journal of Economics, November, 98(4), pp. 659-79.
- Smith, Roy C., and Ingo Walter, 2002, "Rating agencies: Is There an Agency Issue", in Levich, Richard M., Giovanni Majnoni and Carmen Reinhart, eds. Ratings, Rating Agencies and the Global Financial System, Kluwer Academic Publishers, Boston.

- Standard and Poor's, 2008, Standard & Poor's Rating Services U.S. Rating Fees Disclosure, press release (<http://www2.standardandpoors.com/spf/pdf/fixedincome/RatingsFees2008.pdf>).
- Sylla, Richard, 2002, "An historical primer on the business of credit rating", in Levich, Richard M., Giovanni Majnoni and Carmen Reinhart, eds. Ratings, Rating Agencies and the Global Financial System, Kluwer Academic Publishers, Boston.
- Tang, Tony, 2006, "Information Asymmetry and Firms' Credit Market Access: Evidence from Moody's Credit Rating Format Refinement", University of Chicago working paper.
- Weinstein, M., 1977, "The effect of a rating change announcement on bond price", Journal of Financial Economics, 5, 329-350.
- White, Lawrence, 2002, "The Credit Rating Industry: An Industrial Organization Analysis", in Levich, Richard M., Giovanni Majnoni and Carmen Reinhart, eds. Ratings, Rating Agencies and the Global Financial System, Kluwer Academic Publishers, Boston.

**Figure 1.**  
Fitch monthly market share of credit ratings (U.S. issuers)  
12 month moving average 1998 - 2006



**Figure 2.**  
Firm credit ratings distribution: high and low competition in the credit rating industry



**Table 1A: Summary statistics**

	Firm credit rating	Bond credit rating	Fitch market share	Yield-to-maturity	Yield to maturity - matched treasury yield	Leverage	Debt/ EBITDA	Downgrade equity return [-1,1]	Downgrade equity return [-2,2]	Downgrade equity return [-10,10]	Investment-to-junk downgrade return [-2,2]
<b>Mean</b>	18.092	23.080	0.212	6.534	1.381	0.368	3.798	-0.027	-0.027	-0.030	-0.037
<b>Median</b>	18	23	0.225	6.482	1.722	0.343	2.840	-0.011	-0.011	-0.013	-0.015
<b>Standard Deviation</b>	3.930	4.943	0.142	1.690	1.341	0.203	4.073	0.095	0.110	0.133	0.109
<b>Observations</b>	19,756	686,990	429	35,266	33,657	19,756	19,300	1,844	1,837	1,777	221

**Notes:** Each column presents the coefficient estimates from an OLS or logistic specification. Intercepts not reported. The sample period is from 1995 until 2006. The left hand side variable is coded as follows: AAA = 1, AA+ = 2, AA = 3, AA- = 4, A+ = 5, A = 6, A- = 7, BBB+ = 8, BBB = 9, BBB- = 10, BB+ = 11, BB = 12, BB- = 13, B+ = 14, B = 15, B- = 16, CCC = 17, CC = 18, C = 19 and D (default) = 20. Firm characteristics are the measured at the end of the previous fiscal year (using accounting data from Compustat). Leverage is debt over total assets. Downgrade returns refer to cumulative equity returns around a firm downgrade. Investment-to-junk refers to downgrades of firms from investment grade (BBB- and better) to junk status (BB+ and worse).

**Table 1B. Credit ratings**

Rating group	Rating agency		Numerical value assigned*	Category definition**
	Moody's	S&P, Fitch		
Investment Grade	AAA	AAA	28	The obligor's capacity to meet its financial commitment on the obligation is extremely strong.
	Aa	AA	24, 25, 26	The obligor's capacity to meet its financial commitment on the obligation is very strong.
	A	A	21, 22, 23	Somewhat more susceptible to the adverse effects of changes in circumstances and economic conditions than obligations in higher-rated categories. However, the obligor's capacity to meet its financial commitment on the obligation is still strong.
	Baa	BBB	18, 19, 20	Exhibits adequate protection parameters. However, adverse economic conditions or changing circumstances are more likely to lead to a weakened capacity of the obligor to meet its financial commitment on the obligation.
Speculative Grade	Ba	BB	15, 16, 17	Obligations rated 'BB', 'B', 'CCC', 'CC', and 'C' are regarded as having significant speculative characteristics. 'BB' indicates the least degree of speculation and 'C' the highest. While such obligations will likely have some quality and protective characteristics, these may be outweighed by large uncertainties or major exposures to adverse conditions.
	B	B	12, 13, 14	
	Caa	CCC	9, 10, 11	
	Ca	CC	7	
	C	C	4	
Default	D	D	1	An obligation in payment default. The 'D' rating category is used when payments on an obligation are not made on the date due even if the applicable grace period has not expired, unless Standard & Poor's believes that such payments will be made during such grace period. The 'D' rating also will be used upon the filing of a bankruptcy petition or the taking of a similar action if payments on an obligation are jeopardized.

\* Multiple numerical values for a single rating level represent ratings with a + qualifier, no qualifier, and a - qualifier, respectively.

\*\* Source for ratings definitions is Standard & Poor's Ratings Definitions from 17-Mar-2008.

**Table 2. Predicting firm credit ratings with Fitch market share**

Regression model	Dependent Variable: firm credit rating				
	OLS (1)	OLS (2)	OLS (3)	OLS (4)	Ordered Probit (5)
<b>Fitch market share</b>	2.393 *** (1.123)	1.3189 ** (0.564)	0.8040 * (0.417)	0.7795 * (0.430)	0.3615 ** (0.156)
<b>Year Fixed Effects</b>		X	X	X	X
<b>Industry Fixed Effects</b>		X			X
<b>Firm Fixed Effects</b>			X	X	
<b>Firm controls</b>				X	
<b>R<sup>2</sup></b>	0.0057	0.1408	0.8577	0.9032	n/a
<b>N</b>	N = 19,633	N = 19,633	N = 19,633	N = 19,633	N = 19,633

**Notes:** Each column presents the coefficient estimates from an OLS or ordered probit specification. Intercepts not reported. The sample period is from 1995 until 2006. The left hand side variable refers to credit opinion ratings by Standard and Poor's and is coded from 28 (AAA) to 1 (D). See Table 1 for details. Fitch market share is the fraction of bond ratings in an industry-year cell performed by Fitch Ratings. Firm characteristics are the log of sales, log of book value of assets, cash divided by total assets (and it's square), EBITDA divided by total assets (and it's square), cash flow over total assets (and it's square), EBITDA over sales (and it's square), cash flow over sales (and it's square), PPE over total assets (and it's square), interest expense over EBITDA (and it's square), debt over total assets (and it's square), all measured at the end of the previous fiscal year (using accounting data from Compustat). Industries are 2-digit level North American Industry Classifications System (NAICS) industries. The standard errors for the coefficient estimates are in parentheses and are clustered by industry\*year cell.

\* significant at 10%; \*\* significant at 5%; \*\*\* significant at 1%

**Table 3. Predicting bond ratings with Fitch market share**

Regression model	Dependent Variable: bond issue credit rating			
	OLS All ratings	OLS All ratings	OLS All ratings	OLS Excludes NAICS52
	(1)	(2)	(3)	(4)
<b>Fitch market share</b>	0.3685 * (0.1929)	0.4357 ** (0.1853)	0.5129 ** (0.2193)	0.6661 * (0.3438)
<b>Previous rating</b>	0.9860 *** (0.0175)			
<b>Year Fixed Effects</b>	X	X	X	X
<b>Industry Fixed Effects</b>	X	X		
<b>Time to maturity Fixed Effects</b>		X	X	X
<b>Previous rating Fixed Effects</b>		X	X	X
<b>Bond Issue Fixed Effects</b>			X	X
<b>R<sup>2</sup></b>	0.939	0.941	0.944	0.897
<b>N</b>	N = 375,456	N = 368,795	N = 368,795	N = 146,575

**Notes:** Each column presents the coefficient estimates from an OLS or ordered probit specification. Intercepts not reported. The sample period is from 1995 until 2006. The left hand side variable refers to credit opinion ratings by Standard and Poor's and is coded from 28 (AAA) to 1 (D). See Table 1 for details. Fitch market share is the fraction of bond ratings in an industry-year cell performed by Fitch Ratings. Industries are 2-digit level North American Industry Classifications System (NAICS) industries. Previous rating refers to the same bond issue's preceding rating. The standard errors for the coefficient estimates are in parentheses and are clustered by industry\*year cell.

\* significant at 10%; \*\* significant at 5%; \*\*\* significant at 1%



**Table 4. Predicting rating levels with Fitch market share - interactions with leverage**

Regression model	Dependent Variable: firm credit rating			
	OLS (1)	OLS (2)	OLS (3)	OLS (4)
Fitch market share * leverage	6.010 *** (1.645)			
Fitch market share * long term leverage		5.317 *** (1.708)		
Fitch market share * high leverage dummy variable			1.587 ** (0.682)	
Fitch market share * Debt/EBITDA				0.3806 *** (0.129)
Debt/EBITDA				0.3221 *** (0.0483)
Firm controls	X	X	X	X
Industry * Year Fixed Effects	X	X	X	X
R <sup>2</sup>	0.589	0.599	0.588	0.629
N	N = 19,633	N = 19,633	N = 19,633	N = 19,633

**Notes:** Each column presents the coefficient estimates from an OLS or logistic specification. Intercepts not reported. The sample period is from 1995 until 2006. The left hand side variable refers to firm credit opinion ratings by Standard and Poor's and is coded from 28 (AAA) to 1 (D). See Table 1 for details. Fitch market share is the fraction of bond ratings in an industry-year cell performed by Fitch Ratings. Firm characteristics are the log of sales, log of book value of assets, cash divided by total assets (and it's square), EBITDA divided by total assets (and it's square), cash flow over total assets (and it's square), EBITDA over sales (and it's square), cash flow over sales (and it's square), PPE over total assets (and it's square), interest expense over EBITDA (and it's square), debt over total assets (and it's square), all measured at the end of the previous fiscal year (using accounting data from Compustat). Leverage is debt over total assets, , long term leverage is long-term debt over assets, teh high leverage dummy is equal to one if debt over assets is above 0.2324 (the sample median). Industries are 2-digit level North American Industry Classifications System (NAICS) industries. The standard errors for the coefficient estimates are in parentheses and are clustered by industry\*year cell.

\* significant at 10%; \*\* significant at 5%; \*\*\* significant at 1%

**Table 5. Bond yields and ratings - the effect of Fitch market share**

Dependent Variable	Yield to maturity	Yield to maturity	Yield spread	Yield spread (at issue)
Regression model	OLS	OLS	OLS	OLS
	(1)	(2)	(3)	(4)
<b>Credit rating * Fitch market share</b>	0.756 *** (0.054)	0.953 *** (0.204)	0.968 *** (0.214)	0.357 *** (0.131)
<b>Credit rating</b>	-0.567 *** (0.054)	-0.612 *** (0.063)	-0.620 *** (0.063)	-0.276 *** (0.030)
<b>Fitch market share</b>	-14.283 ** (3.699)			
<b>Log of time to maturity</b>	-0.475 * (0.275)	-0.484 * (0.273)	-1.204 *** (0.165)	0.487 *** (0.062)
<b>Log of time to maturity, squared</b>	0.212 *** (0.048)	0.222 *** (0.048)	0.268 *** (0.035)	-0.079 *** (0.017)
<b>Log of offering amount</b>	-0.009 (0.116)	-0.099 (0.126)	-0.302 *** (0.108)	-0.453 (0.307)
<b>Log of offering amount, squared</b>	-0.007 (0.005)	-0.003 (0.005)	0.006 (0.004)	-0.017 (0.013)
<b>Date Fixed Effects (Month - Year)</b>	X	X	X	X
<b>Industry Fixed Effects</b>	X			
<b>Year - Industry Fixed Effects</b>		X	X	X
<b>R<sup>2</sup></b>	0.567	0.594	0.503	0.521
<b>N</b>	N = 113,125	N = 113,125	N = 110,965	N = 9,161

**Notes:** Each column presents the coefficient estimates from an OLS regression. Intercepts not reported. Each observation is the yield to maturity of a bond in one transaction. The sample period is from 1995 until 2006. The left hand side variable in the column one to four is the yield to maturity implied by the price in a bond trade reported by FISD. In column five, the dependent variable is the yield to maturity minus the yield to maturity of the government bond with the closest maturity. Credit ratings are bond credit rating issued by Standard and Poors and Moody's (reported by FISD), and represent the latest preceding the transaction (if several were issued simultaneously, we use the average). Fitch market share is the fraction of bond ratings in an industry-year cell issued by Fitch Ratings. Industries are 2-digit level North American Industry Classifications System (NAICS) industries. Bonds are excluded if they have non-standard features (see text for details) or negative yields or yields above 20%. The standard errors for the coefficient estimates are in parentheses and are clustered by industry\*year cell.

\* significant at 10%; \*\* significant at 5%; \*\*\* significant at 1%

**Table 6. Announcement returns around Firm credit rating downgrades - the effect of Fitch market share**

Regression model [t1,t2]	Dependent Variable: equity return [t1,t2]					
	OLS [-1,1]	OLS [-2,2]	OLS [-10,1]	OLS [-2,2]	OLS [-2,2]	OLS [-2,2]
	(1)	(2)	(3)	(4)	(5)	(6)
<b>Fitch market share</b>	-0.046 (0.037)	-0.093 * (0.053)	-0.112 ** (0.057)	-0.094 * (0.050)	-0.290 *** (0.090)	-0.307 *** (0.097)
<b>Volatility</b>				-1.059 (1.303)		-10.820 (6.858)
<b>Volatility squared</b>				-10.092 (28.57)		305.4 * (173.51)
<b>Volatility cubed</b>				122.7 (158.7)		-2695.8 ** (1272.7)
<b>Industry Fixed Effects</b>	X	X	X	X	X	X
<b>Year Fixed Effects</b>	X	X	X	X	X	X
<b>R<sup>2</sup></b>	0.043	0.053	0.038	0.109	0.287	0.364
<b>N</b>	N = 1,585	N = 1,580	N = 1,533	N = 1,552	N = 182	N = 179

**Notes:** Each column presents the coefficient estimates from an OLS regression. Intercepts not reported. Returns are for firm equity, as reported in the CRSP database. The sample period is from 1996 until 2006. The left hand side variable is the cumulative equity return from time t1 to t2 (where zero represents the day of the downgrade return around a rating downgrade), calculated from CRSP data and net of the value-weighted market return. Observations with event returns larger than 50 percent or lower than minus 50 percent are excluded. Volatility is the standard deviation of daily stock returns in the preceding 120 trading days. Fitch market share is the fraction of bond ratings in an industry-year cell performed by Fitch Ratings. Industries are 2-digit level North American Industry Classifications System (NAICS) industries. The standard errors for the coefficient estimates are in parentheses and are clustered by industry\*year cell.

\* significant at 10%; \*\* significant at 5%; \*\*\* significant at 1%

**Table 7. Fitch market share and segment characteristics**

<b>Dependent Variable: Fitch's market share</b>					
<b>Dependent variable:</b>	# Ratings issued	Industry debt (\$)	Change in ind. debt (\$)	Industry profitability	Ind. profitability (median)
<b>Timing of Fitch's market share</b>	(1)	(2)	(3)	(4)	(5)
<b>Lead (t+2)</b>	0.1831 (0.4055)	-0.1079 (0.1937)	-0.0885 (0.2039)	-0.0108 (0.0272)	0.0085 (0.0156)
<b>Lead (t+1)</b>	0.0194 (0.4210)	0.0824 (0.1913)	0.2870 (0.1794)	-0.0167 (0.0248)	0.0039 (0.0155)
<b>Simultaneous (t)</b>	0.1474 (0.4151)	-0.0617 (0.1739)	-0.1813 (0.1541)	-0.0086 (0.0251)	0.0112 (0.0158)
<b>Lag (t+1)</b>	0.2968 (0.3367)	-0.2615 (0.1673)	-0.1362 (0.1584)	0.0151 (0.0245)	-0.0008 (0.0149)
<b>Lag (t+2)</b>	0.2015 (0.3190)	-0.3075 * (0.1768)	0.1018 (0.1577)	0.0213 (0.0237)	0.0221 (0.0165)
<b>Industry Fixed Effects</b>	X	X	X	X	X
<b>Year Fixed Effects</b>	X	X	X	X	X
<b>N</b>	$N \leq 266$	$N \leq 266$	$N \leq 266$	$N \leq 266$	$N \leq 266$

**Notes:** Each coefficient estimate refers to one OLS specification (different rows represent regressions which differ only in the timing of the independent variable). Each regression includes year and industry fixed effects. For each regression, the coefficient estimate for Fitch's market share is reported. The sample period is from 1995 until 2006. Number of ratings issued is the log of the aggregate number of credit ratings issued for bonds in an industry. Industry debt is the log of the total amount of outstanding debt of Compustat firms in an industry. The change in industry debt is the the amount of industry debt minus it's previous value in the same industry. Industry profitability is an asset-weighted average of the ratio of EBITDA to assets in Compustat firms in the industry and Industry profitability (median) is the median EBITDA-asset ratio across all Compustat firms in the industry. ZZZ Fitch market share is the fraction of bond ratings in an industry-year cell performed by Fitch Ratings. The number of observations is 266 or fewer (some observations are lost due to lags). The standard errors for the coefficient estimates are in parentheses and are heteroskedasticity-robust.

\* significant at 10%; \*\* significant at 5%; \*\*\* significant at 1%